// NO.22-004 | 02/2022

DISCUSSION PAPER

// LUKAS RIEDEL AND HOLGER STICHNOTH

Government Expenditure in the DINA Framework: Allocation Methods and Consequences for Post-tax Income Inequality





Government Expenditure in the DINA Framework: Allocation Methods and Consequences for Post-Tax Income Inequality*

Lukas Riedel

ZEW Mannheim and University of Heidelberg

Holger Stichnoth[†]

ZEW Mannheim and University of Strasbourg

February 8, 2022

Abstract

Constructing measures of post-tax income inequality that are consistent with national accounts requires the allocation of the entirety of government expenditure to individuals. About half of government expenditure in the United States takes the form of in-kind collective expenditure (e.g., education, defense, infrastructure). The dominant assumption in the literature is to allocate this expenditure proportionally to post-tax cash income. We show that the gap in post-tax income shares between the Top 10% and Bottom 50% in the United States is reduced by half (from about 20 to 10 percentage points in recent years) when this assumption is replaced by a lump-sum allocation. We further provide direct evidence on how a substantial part of collective expenditure is actually distributed. When adopting the cross-sectional perspective of the DINA approach, we find that public education spending goes disproportionately to the bottom half of the income distribution. A lump-sum allocation provides a good approximation. Moving beyond the cross-section, we find that public education expenditure is positively correlated with both lifetime earnings and parents' socio-economic status.

Keywords: inequality, redistribution, education, in-kind transfers

JEL Codes: D31, H41, H52, I24

^{*}We thank Christina Gathmann, Valentina Melentyeva, Sebastian Siegloch, and Michaela Slotwinski as well as seminar participants at Mannheim, Strasbourg, and ZEW, and conference participants at ECINEQ, IIPF, and Verein für Socialpolitik for valuable comments and suggestions. Hanne Albig and Lena Göhringer provided excellent research assistance.

[†]Corresponding author. ZEW Mannheim, Research Department Social Policy and Redistribution. L7, 1 68161 Mannheim Germany. holger.stichnoth@zew.de.

1 Introduction

The United States and many other countries have seen an increase in income inequality in recent decades that has received attention from academic researchers and the general public alike. The measurement of income inequality has traditionally relied on micro data from surveys or administrative tax records. These data, however, capture only about 60% of macro totals from national accounts, so a substantial share of national income has been missing from the debate about inequality. In an important contribution, Piketty, Saez, and Zucman (2018) propose a method for constructing distributional national accounts (DINA) that measure how the entire national income is distributed among individuals. When computing post-tax income, this requires the allocation of the entirety of government expenditure to individuals. In recent years, about half of government expenditure in the United States has taken the form of in-kind consumption expenditure (e.g., education, defense, infrastructure); depending on the year, this represents between 16% and 20% of national income. Piketty, Saez, and Zucman assume that this collective in-kind expenditure is distributed proportionally to post-tax disposable income. This means that, by construction, an important part of national income is assumed to be distributionally neutral. The recent version of the DINA Guidelines (Alvaredo et al. 2020) explicitly recognizes the difficulty surrounding the allocation of government consumption expenditure, calling it "approximate and exploratory."

The present paper makes three contributions to the debate about the measurement of post-tax income inequality. We first revisit the DINA study by Piketty, Saez, and Zucman (2018) for the United States and show that their assumption regarding government consumption expenditure implies high per-capita levels of public expenditure for the very rich. Moreover, and more importantly, the assumption has a substantial effect on the level of post-tax income inequality. When we replace Piketty, Saez, and Zucman's proportionality assumption with a lump-sum allocation, the Top 10% share of national income decreases by about 5 percentage points, while the share of the Bottom 50% increases by roughly the same amount. As a result, the gap between the income shares of the Top 10% and the Bottom 50% is reduced by half, from about 20 to 10 percentage points in the most recent years. The trend is hardly affected, however.

Our second and main contribution is to provide direct evidence on how an important fraction of government consumption expenditure is actually distributed. We focus on public spending on education, which makes up about 30% of collective expenditure and 5% of national income in most OECD countries, and is much easier to assign indi-

¹The assumptions regarding the allocation of government expenditure also matter further up in the income distribution. Piketty, Saez, and Zucman (2018) find that 15.7% of post-tax income goes to the Top 1% in 2014. Auten and Splinter (2019) report a share of 8.6%. In their decomposition of this difference, 1.3 percentage points are due to a different allocation of government consumption. While Piketty, Saez, and Zucman assign all of it proportional to after-tax income, Auten and Splinter allocate 50% of it on a per-capita basis.

vidually than defense or infrastructure expenditure. Our data for the United States are from the 2017 wave of the American Community Survey (ACS). In addition to the large sample size (about 3.2 M individuals in 1.4 M households), the ACS has the advantage that participants are legally obligated to answer the survey questions. The ACS has information on whether household members are currently in education, and, importantly for our purpose, distinguishes between public and private institutions. Finally, the ACS includes individuals in group quarters, which is key for measuring public expenditure that goes to college students who no longer live with their parents. Annual public expenditure per student (net of tuition fees) at different levels of education is taken from the OECD.

We find that, for education at least, public expenditure is not proportional to income. On the contrary, average public education spending is highest in the poorest decile and lowest in the richest decile. The differences are not great, however, so a lump-sum allocation provides a good approximation, at least when income is measured using the equal-split assumption of the DINA framework (i.e., household income is divided by the number of adults aged 20 and above). When equivalized household income is used instead, the negative income gradient is steeper and the approximation is less accurate.

These results are strongly driven by age effects. The most striking case are college students who no longer live with their parents. They receive substantial public expenditure while having low current income. But public spending at other levels (pre-primary, primary, secondary) also has an age component, as parents with kindergarten- or school-age children are typically still below the peak of their age-income profiles.

In our third contribution, we examine two justifications for an allocation of collective expenditure proportionally to income that have been proposed in the DINA literature. Piketty, Saez, and Zucman (2017) argue for a proportional allocation by pointing to the positive correlation between public education spending and lifetime earnings. Using the American Community Survey and proxying for lifetime earnings using earnings at age 40–45 (where the rank correlation with lifetime earnings is maximal), we quantify this argument by showing that the 10% of individuals with the highest earnings have received average public education spending of \$335 K, about 1.4 times the amount of the bottom 50% (\$234 K). The allocation is still not proportional to earnings, however; proportionality would require a factor of about 14. More importantly, adjusting for age effects in public education spending, but not in earnings, capital income, or certain cash transfers, would be inconsistent with the DINA framework, which so far has adopted a

²Two caveats apply. First, the American Community Survey does not provide the comprehensive income measure that is the *raison d'être* of the DINA approach. However, additional income components such as imputed rents and especially undistributed profits are concentrated among the higher deciles, thus leading to an even greater departure from proportionality. The second caveat is that the ACS provides pre-tax income, while Piketty, Saez, and Zucman (2018) assume that government consumption expenditure is proportional to post-tax income. However, when we simulate post-tax income based on the ACS pre-tax measure and the NBER's TAXSIM model (Feenberg and Coutts 1993), we still clearly reject the proportionality assumption.

strictly cross-sectional perspective.

The DINA Guidelines (Alvaredo et al. 2020) argue that a lump-sum allocation would overestimate the extent of redistribution because of the unequal access to education observed in most countries. While the American Community Survey does not allow us to address this point, we leverage the biographical information in the German Socio-Economic Panel to show that more public education spending indeed goes to children of parents with higher education or greater occupational prestige. The difference between individuals from the most and the least privileged background with respect to parents' education is \$57 K on the father's side and \$49 K on the mother's side. For parents' occupational prestige (measured using the occupation they held when the respondents were 15 years old), the differences are slightly less pronounced, but still sizeable. However, while these inter-generational patterns are arguably more important than the cross-sectional results for the distributional debate, they again do not provide the right empirical basis for an allocation of government expenditure in the cross-section.

Related literature Following the paper by Piketty, Saez, and Zucman (2018) for the United States, the DINA approach has been applied to other countries. Garbinti, Goupille-Lebret, and Piketty (2018) study pre-tax income inequality in France using a DINA approach, and Bozio et al. (2018) extend this to post-tax income. Using a simplified approach, Blanchet, Chancel, and Gethin (2019) create distributional national accounts for the member countries of the European Union. Other application of the DINA framework are from Austria (Jestl and List 2020), China (Piketty, Yang, and Zucman 2019), Germany (Bach, Bartels, and Neef 2021), and Sweden (Hammar et al. 2020). In a related effort, the OECD and Eurostat set up an expert group to disaggregate the household sector in the system of national accounts; see Zwijnenburg (2019) for a comparison with the DINA approach.

Our paper contributes to the discussion about methodological issues in the measurement of income inequality in the DINA framework and beyond. Note that we focus exclusively on the effect of government (in-kind) consumption expenditure and remain silent on the debate about issues in the measurement of pre-tax income, such as the allocation of business profits or untaxed pension income (Auten and Splinter 2019; Saez and Zucman 2020).³

There is an older literature on the distribution of public (in-kind) expenditure (Callan, Smeeding, and Tsakloglou 2008; Garfinkel, Rainwater, and Smeeding 2006; Gillespie 1965; Horton and Reed 2010; Marical et al. 2006; Musgrave, Case, and Leonard 1974; O'Dea and Preston 2012; O'Higgins and P. Ruggles 1981; P. Ruggles and O'Higgins 1981; Smeeding 1977; Smeeding et al. 1993; Verbist, Förster, and Vaalavou 2012; Zwijnenburg, Bournot,

³There is a also a debate about measurement issues regarding wealth inequality, see Saez and Zucman (2016), Smith et al. (2019) and Saez and Zucman (2020).

and Giovannelli 2017). Several of these studies also study public education spending. The patterns found in these studies are consistent with our results. In particular, none of the studies find that the allocation of public education spending is proportional to cash income. We contribute to this literature by using a much larger dataset that distinguishes between public and private education as well as different levels of education (pre-primary, primary, secondary, tertiary) and that includes students in group quarters, which is important for the allocation of public spending on tertiary education. We also contribute by linking our findings to the DINA literature. In particular, we break down education spending by individualized income for adults age 20 and above, using the "equal-split" approach of Piketty, Saez, and Zucman (2018). The older studies used equivalized household income instead, which we include as a robustness check. Finally, while the earlier studies examine public spending in the cross-section, we additionally distinguish by lifetime earnings and by the socio-economic status of the parents.

This earlier literature has raised the important question of whether government in-kind expenditure should be measured at cost or should rather measure the increase in individual welfare that results from the expenditure (see O'Dea and Preston 2012, on this and other methodological issues). With an assignment based on cost, inefficiencies in the provision of public services show up as income, and there is no accounting for different needs of individuals. However, attempts to measure welfare instead of income or to account for different needs by adjusting equivalence scales (Aaberge, Bhuller, et al. 2010; Aaberge, Eika, et al. 2019; Aaberge, Langørgen, and Lindgren 2013; Paulus, Sutherland, and Tsakloglou 2010) depart from the DINA framework, which – following the practice in national accounts – measures government expenditure on a cost basis. Moreover, we see the issue of valuation as orthogonal to the question of correctly determining who receives the public expenditure in the first place.

The remainder of this paper is organized as follows. Section 2 revisits the paper by Piketty, Saez, and Zucman (2018). Based on their publicly available data, we contrast two assumptions (proportionality to post-tax cash income versus lump-sum) regarding the allocation of government consumption expenditure, and show that these assumptions have a substantial effect on the level, but not the trend, of post-tax income inequality. Section 3 describes the data and methods we use in our empirical study of how public education spending is actually distributed. Section 4 presents our results. We focus on the distribution in the cross-section, which is the perspective that has been adopted in the DINA literature, but also report the distribution by life-time earnings (proxied for by earnings at age 40–45). Finally, in a supplementary analysis based on German data, we study how public education expenditure varies by parents' socio-economic status. Section 5 concludes.

2 Revisiting Piketty, Saez, and Zucman (2018)

2.1 Overview

In an important methodological contribution, Piketty, Saez, and Zucman (2018) create distributional national accounts that make income measures from tax and survey data consistent with the macro totals published in national accounts. They study pre-tax and post-tax income inequality in the United States for the years 1913–2014 and document a massive increase for both types of inequality since 1980.

In the first part of our article, we argue that their findings regardings the *level* of post-tax inequality are very sensitive to their assumption regarding the allocation of government consumption expenditure. Based on their publicly available data, we show that with a different assumption—a lump-sum allocation of collective expenditure—, the gap in the shares of post-tax national income accruing to the Bottom 50% and the Top 10% is reduced by half in recent years, from 20 to 10 percentage points.⁴

2.2 Post-tax Income Inequality and Government Expenditure

Figure 1 summarizes the distribution of U.S. national income in 2014, the most recent year in their study. Piketty, Saez, and Zucman allocate all items of national income to adults age 20 and above. In couples, the income is assumed to be split equally. The mean value of national income by adult in 2014 is \$65 K. By construction, the mean is the same for pre-tax and post-tax income, which are alternative ways of allocating the same total national income.⁵ Pre-tax income is distributed very unequally: the 10% of adults with the highest pre-tax income receive 47% of the total, while the Bottom 50% receive only 13%. This translates into an average pre-tax income of about \$300 K among the Top 10% (47/10 times the mean income of \$65 K), compared with \$16 K for the bottom half of the pre-tax income distribution. The Middle 40% receive almost exactly their population share of 40%, and accordingly have an average pre-tax income close to the overall mean.

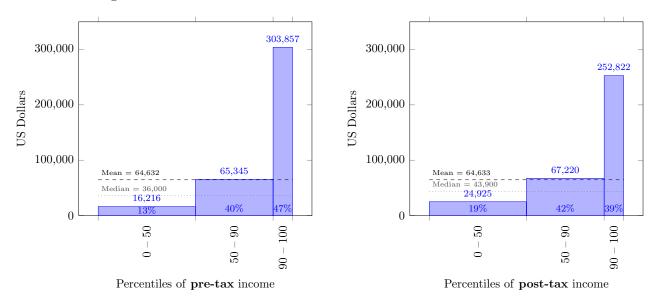
Income after taxes and transfers is only slightly less unequally distributed. The share of the Top 10% decreases from 47% to 39%, while the shares of the Middle 40% and the Bottom 50% increase by 2 and 6 percentage points, respectively.

Figure 2 shows how post-tax income is divided between two broad categories – income net of taxes on the one hand and transfers on the other – and how each is divided among the Top 10%, Middle 40%, and Bottom 50%. Overall, 66.5% of U.S. national income in

⁴The results, code, and most of the micro data are available at http://gabriel-zucman.eu/usdina/. We use the November 2017 vintage, which corresponds to the published version (Piketty, Saez, and Zucman 2018). The series have since been updated to more recent years, improved, and revised (to incorporate changes in the underlying National Accounts data). These changes are documented in https://gabriel-zucman.eu/files/PSZUpdates.pdf. The part of the analysis that we focus on in this article – the allocation of government consumption expenditure – has not been affected by the updates.

⁵The small difference in Figure 1—63,632 vs. 64.633—is due to rounding.

Figure 1: Distribution of Pre-tax and Post-tax National Income

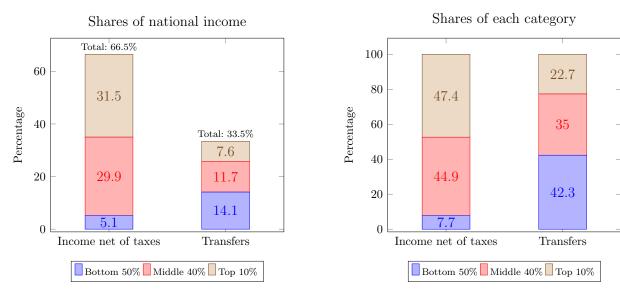


Notes: The figure shows how national income is distributed among adults aged 20 and above in the United States in 2014. The figure depicts the income shares of the Bottom 50%, Middle 40%, and Top 10%, both for total pre-tax income (left panel) and for post-tax income (right panel), as well as the overall mean and median and the mean within each group. Source: Own calculations based on Piketty, Saez, and Zucman (2018). Pre-tax income: Appendix Tables II-B1, II-B3, II-B13. Post-tax income: Appendix Tables II-C1, II-C3, II-C13.

2014 corresponds to income net of taxes, while the remaining 33.5% are transfers. The share of national income that goes to the Bottom 50% is made up of 5.1% of income net of taxes and 14.1% of transfers, yielding a total of 19%. For the two other groups, post-tax income is mostly income net of taxes, but transfers play a role as well. In fact, the Top 10% receive more than twice their population share in terms of transfers (22.7%), while the Bottom 50% receive less than half of all transfers (42.1%).

This surprising result is explained by the way in which government transfers are allocated to individuals in Piketty, Saez, and Zucman (2018)'s analysis. Figure 3 breaks down government transfers into several underlying categories. Overall, these transfers amount to \$5.072 B or 33.5% of total national income (\$15.154 B) in 2014. Piketty, Saez, and Zucman treat about half of these transfers (\$2.515 B) as individualized. This category in turn can be divided into cash transfers and in-kind transfers. The cash transfers are Social Security pension and non-pension (disability insurance, unemployment insurance), social assistance benefits in cash (refundable tax credits, veterans' benefits, workers' compensation, food stamps, supplemental security income, TANF/AFDC, and some smaller programs). These are assigned based on rules and on the recipient status observed in the Current Population Survey (CPS). Individualized in-kind transfers are mostly Medicare (assigned based on rules: age or receipt of disability insurance) and Medicaid (assigned based on the CPS). Note that some of these transfers (pension benefits, disability, and unemployment insurance) are already included in pre-tax income and are thus not counted

Figure 2: Decomposition of Post-tax Income



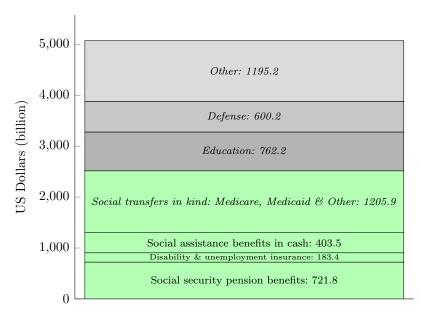
Notes: The figure shows a decomposition of post-tax income into income net of taxes and transfers. The left panel shows that income net of taxes makes up 66.5% of national income, while transfers make up the remaining 33.5%. The right panel decomposes both categories of national income. The figure shows, for example, that the transfers that accrue to the Bottom 50% represent 14.1% of national income and 42.3% of all transfers. Source: Own calculations for the United States in 2014 based on Piketty, Saez, and Zucman (2018), Appendix Table II-C2.

towards as government redistribution in the definition of Piketty, Saez, and Zucman, which is limited to the difference between pre-tax and post-tax income.

The other half of government transfers (\$2.558 B) represents in-kind transfers in three domains: education, defense, and a catch-all other category, which includes roads, public transportation and more generally the physical as well as legal and administrative infrastructure. These are items of government consumption expenditure. They represent goods and services and not a cash flow from the government to individuals. In accordance with the practice of national accounting, they are valued at the monetary cost of providing them (net of fees for their use), as opposed to the monetary equivalent of the benefit that individuals attach to them, which is much more difficult to measure. Citing the difficulty of observing who receives these goods and services (except for education), Piketty, Saez, and Zucman opt to allocate all of them proportionally to post-tax disposable income, which is pre-tax income minus taxes plus individualized monetary transfers.

This choice makes half of government spending distributionally neutral by assumption, and implies extremely unequal amounts of government consumption expenditure per capita (Figure 4a). As the 50% of adults with the lowest post-tax disposable income receive 18.0% of the total, they get assigned the same share of government consumption expenditure, which corresponds to less than \$4 K per person and year. By contrast, each adult in the Top 10% is assumed to receive \$45 K per year in terms of public spending on education, defense, public transportation, roads, and other infrastructure, despite more

Figure 3: Categories of Government Expenditure



Notes: The figures shows the different categories that make up total government expenditure. Individualized transfers are shown in green, collective expenditure in gray. Transfers in kind are italicized, the remaining items are cash transfers. Social assistance in cash comprises refundable tax credits, SNAP, SSI, TANF/AFDC, and various smaller programs. Source: Own calculations for the United States in 2014 based on Piketty, Saez, and Zucman (2018), Appendix Table I-SA11.

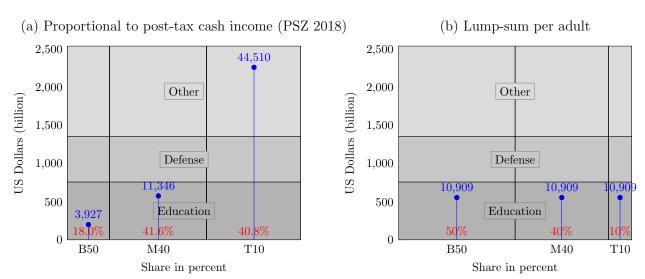
frequently using private-sector alternatives, at least for education and transportation. At the very top per capita values are even higher. The 0.01% with the highest incomes each receive more than \$4 M per year.

2.3 Consequences for Post-tax Income Shares

Levels With a lump-sum allocation, each adult gets assigned the same value of annual government consumption expenditure of \$11 K (see Figure 4b). This assumption leads to a substantial change in the level of post-tax inequality. With a lump-sum allocation, the gap in the post-tax income shares between the Bottom 50% and the Top 10% is reduced by half in 2014. When each adults is allocated the same amount of government consumption expenditure, the share of the Bottom 50% is higher by about 5 percentage points, and the share of the Top 10% is reduced by about the same magnitude compared with an allocation that is proportional to post-tax disposable income (see Figure 11 in the Appendix). As a result, the gap in the income shares of the two groups is reduced from about 20 to 10 percentage points. The share of the Middle 40% is almost unaffected. The effect of the allocation rule on the income shares is of the same order of magnitude as in the study by Blanchet, Chancel, and Gethin (2019) for a number of European countries.

Trends The sensitivity of the level of post-tax income inequality to the assumptions regarding the allocation of government consumption has not always been highlighted

Figure 4: Comparison of Assumptions About Collective Expenditure



Notes: The figure contrasts Piketty, Saez, and Zucman (2018)'s assumption regarding the allocation of collective expenditure with the alternative of a lump-sum allocation. Piketty, Saez, and Zucman allocate collective expenditure proportionally to post-tax cash income (left panel). With this assumption, the Bottom 50% of the post-tax income distribution receive 18.0% of collective expenditure, while the Middle 40% receive 41.6%, and the Top 10% 40.8%. This implies a per-capita value of \$3,927 in the bottom half of the distribution, compared with \$44,510 among the Top 10%. The right panel shows an alternative assumption in which each adult receives the same share of collective expenditure, which corresponds to a per-capita value of \$10,909. With this assumption, the share of collective expenditure that goes to the three groups is equal to their population share. Source: Own calculations for the United States in 2014 based on Piketty, Saez, and Zucman (2018), Appendix Tables I-SA11, II-C1b.

enough in the DINA literature⁶ and motivates our analysis of how this expenditure (or parts thereof) is actually distributed. However, before we turn to the empirical part of our paper, we would like to emphasize that the *trend* in post-tax income inequality is probably much less affected by the allocation. In particular, the key finding of Piketty, Saez, and Zucman, namely the sharp increase not only in pre-tax, but also post-tax inequality over the past four decades or so, also holds with a lump-sum allocation of government consumption expenditure.

As Figure 5 shows, replacing the proportional allocation with a lump-sum allocaton leads to a parallel shift in the series for the national income shares of the Bottom 50% and the Top 10%. With the lump-sum allocation, the series intersect both in the mid-1960s and the mid-1980s. However, given that the population shares of the two groups differ, an identical share of national income means that the average post-tax income of the Top 10% is five times larger than for the Bottom 50%. In 2014, the ratio of average incomes is 10.1 with a proportional allocation and 6.9 with a lump-sum allocation.

There are two reasons for the parallel shift. First, the share of collective expenditure in national income has been fairly stable between 15 and 20% over the period considered here. Second, while the income shares based on a proportional allocation merely reflect the trends observed for post-tax disposable income, the series for the lump-sum allocation is based on population shares that are time-constant by construction (Top 10%, Middle 40%, Bottom 50%) and thus cannot capture any real movements in the allocation of government consumption expenditure either. We do not believe that the allocation of these expenditure items has changed so much since 1980 that the divergence in post-tax income shares documented by Piketty, Saez, and Zucman would be altered substantially.

3 Methods and Data

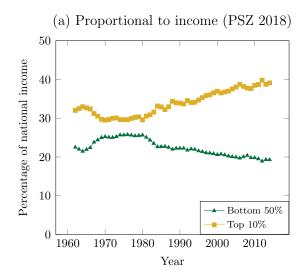
3.1 Overview

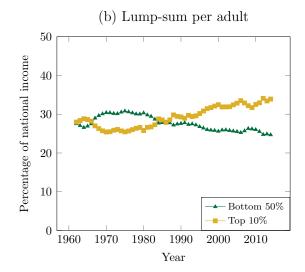
The previous section has shown that the level of post-tax income inequality is sensitive to the assumption about how government consumption expenditure is allocated. In the remainder of this paper, we provide direct evidence on how an important part of this expenditure is actually distributed. We focus on public spending on education, which makes up about 5% of national income in the US and in most OECD countries and is much easier to assign individually than defense or infrastructure expenditure.

Our method for allocating public education expenditure is straightforward. We use a micro dataset—the American Community Survey 2017—that allows us to observe the

 $^{^6}$ Piketty, Saez, and Zucman (2018) do run a robustness check (to which we will return below) in which they assign public education spending not proportional to post-tax income, but as a function of the number of children. They only report the consequences for the average income of the Bottom 50% and not the shift in the income shares of the three groups.

Figure 5: Effect of the Assumptions on Post-tax Income Shares, 1962–2014





Notes: The figure shows how the assumption regarding the allocation of collective expenditure affects the distribution of post-tax income in the United States over the years 1962–2014. Each panel shows the share of the Bottom 50% and the Top 10%. The left panel is for the assumption adopted by Piketty, Saez, and Zucman (2018), i.e. an allocation of collective expenditure that is proportional to post-tax cash income. The right panel shows the income shares that result from assuming a lump-sum allocation. Source: Own calculations based on Piketty, Saez, and Zucman (2018), Appendix Tables I-SA11, II-C1b, II-C2, II-C3b.

income of the household and that has information on who in the household currently attends a *public* educational institution, distinguishing pre-primary, primary, secondary, and tertiary education. We then multiply the number of students per household with the average public expenditure for students of the respective education level, which we take from the OECD's "Education at a Glance" database.

Following the DINA framework, our main analysis is cross-sectional, i.e. we study the distribution of public education expenditure by current income. In addition, we analyze public education expenditure by lifetime earnings, proxied for by earnings at age 40–45. However, based on another dataset—the German Socio Economic Panel, SOEP—we also adopt an intergenerational perspective and document how the expenditure differs by parents' education and occupational prestige.

3.2 American Community Survey

Our main source of individual-level microdata is the American Community Survey (ACS). The ACS is conducted by the United States Census Bureau to collect information similar to the decennial census. Our data for the year 2017 is from the public use file of the ACS provided by IPUMS USA (S. Ruggles et al. 2020). It provides information on around 3.2 M individuals in 1.4 M households. In addition to the large sample size, the ACS has the advantage that—unlike in other datasets such as the Current Population Survey—respondents are legally obligated to answer the survey questions.

Enrollment The ACS has information on whether household members are currently enrolled in an educational institution, and, importantly for our purpose, distinguishes between public and private institutions. Moreover, the ACS includes individuals in group quarters including college dormitories, which is key for measuring public expenditure that goes to college students who no longer live with their parents.

The ACS provides a very accurate picture of the number of individuals enrolled in the education system (Figure 13 in the Appendix). For public institutions at the preprimary, primary, and secondary levels in 2017, our own calculations based on the ACS result in 51.4 M students. The OECD (OECD Statistics 2020) and the National Center for Education Statistics (De Brey et al. 2021) report values of 50.6 M and 50.7 M, respectively. At the tertiary level, our ACS number is 16.8 M, which is a little higher than the value of 14.6 M reported by the OECD and the NCES. For completeness, Figure 13 also shows the number of students in private education, although we do not include these students when allocating public education expenditure. Private education is empirically relevant only at the pre-primary level (kindergarten) and then again at the tertiary level. Our ACS numbers are again close to the OECD values, while the numbers reported by the NCES are slightly lower.

Income concept Income is measured in the ACS as an aggregate of personal income from different sources—including wages (including bonuses), income from self-employment and businesses, pensions, capital income, and social security payments—of all household members above the age of 15.8 For individuals in group quarters, such as students in college dormitories, the concept of household income does not apply and only personal income is reported. The period of reference for the income measurement are the previous twelve months. Note that, as the ACS is administered throughout the year, this means that the income in most cases does not correspond to a calendar year. Also, despite the legal obligation to answer the survey, some of the individual income components are actually imputed by the data provider. In a robustness check, we drop all households in which more than half of household income is based on an imputation.

Regarding the comparison with the DINA approach, two additional caveats are in

⁷In a robustness check, we scale down the ACS numbers accordingly.

⁸ "Personal income, or 'money income,' as per the Census Bureau, is the income received on a regular basis (exclusive of certain money receipts such as capital gains and lump-sum payments) before payments for personal income taxes, Social Security and Medicare taxes, union dues, etc. It includes income received from wages, salary, commissions, bonuses, and tips; self-employment income from own nonfarm or farm businesses, including proprietorships and partnerships; interest, dividends, net rental income, royalty income, or income from estates and trusts; Social Security or Railroad Retirement income; Supplemental Security Income (SSI); any cash public assistance or welfare payments from the state or local welfare office; retirement, survivor, or disability benefits; and any other sources of income received regularly such as Veterans' (VA) payments, unemployment and/or worker's compensation, child support, and alimony." (https://www.pewresearch.org/social-trends/2018/07/12/methodology-15/). The income components such as wage or business income are top-coded at the 99.5th percentile of the respective federal state. Higher values are coded as the state-specific average of all values above the threshold.

order. First, while the American Community Survey provides a fairly comprehensive measure of income, it falls short of the DINA approach, in which pre-tax income sums up to the whole of national income. However, additional income components such as imputed rents and especially undistributed profits are concentrated among the higher deciles, and including them would likely lead to an even greater departure from proportionality than what we find below for our income measure.

The second caveat is that the ACS provides pre-tax income, while Piketty, Saez, and Zucman (2018) assume proportionality of government consumption expenditure to post-tax disposable income. However, when we simulate post-tax income using NBER TAXSIM, we still clearly reject the proportionality assumption.⁹

Unit of measurement In our main specification, we follow Piketty, Saez, and Zucman (2018) and the other Dina studies and measure income at the level of adult individuals aged 20 and above. For couples, we apply an equal-split rule, i.e. each adults gets assigned the same share of household income, while children are disregarded. We apply this rule also in cases in which there are more than two adults in the household (e.g., children over 20 or other relatives).

Another approach in the measurement of inequality takes the household as the unit of measurement and accounts for differences in household size and age composition through equivalence scales. This is the approach that is used in the older literature on the measurement of in-kind transfers (see Section 1), and we report results in this tradition as well. When doing so, we apply the commonly used modified OECD equivalence scale, which assigns a value of 1 to the first adult in the household, of 0.5 to each additional household member aged 14 and above, and of 0.3 to each child below the age of 14.

Summary statistics Table 1, panel A, shows summary statistics for our main sample of adults age 20 and above. These represent about 2.4 M, or 75% of the 3.2 M individuals—adults and children—in the ACS. The average age in our sample is 48.4 years. Age is highest in the second and tenth deciles and first falls and then rises in the deciles in between. The first decile is not part of this U shape and stands out for having the lowest average age.

The first and the tenth deciles have the smallest household size on average (2.70 and 2.71). In between, the pattern is an inverted U, with a maximum of around 3.20 in deciles three and four. The differences in household size are mostly driven by the intensive margin. With the exception of the first decile, where the share of adults with children is not only much lower than the average but also noticeably less than in the decile just

⁹TAXSIM (Feenberg and Coutts 1993) simulates the tax liability for federal, state, and payroll taxes. We use TAXSIM version 32 (https://users.nber.org/~taxsim/taxsim32/). The simulations are for tax units, which we identify in our ACS household sample following the procedure outlined by Samwick (2013). We assume that all married couples file jointly.

above, there is little variation in this share across the other deciles, with a slight increase towards the upper range of the income distribution. The age of the youngest child likewise increases with income. The same patterns with respect to age and household composition also hold when grouping individuals based on their post-tax disposable income (panel B).

Turning to the income measures themselves, the mean value of pre-tax income in our ACS sample for the year 2017 is \$43.9 K per adult, and the median is \$32.2 K. Our median is reasonably close to the value of \$36.0 K reported by Piketty, Saez, and Zucman (2018) for 2014, while our mean is much lower than the \$64.6 K that they find using their more comprehensive measure of income (see Figure 1 above). For the mean, we can also compare the values by decile. The difference is mostly driven by the richest decile, where the average reported by Piketty, Saez, and Zucman is almost twice as high as the one we compute based on the income concept from the American Community Survey, which does not include imputed rents and undistributed profits (and additionally is right-censored). If these were to be included, we would find an even bigger departure from a proportional allocation than we actually do.

Note that the income values in the first decile are very low, with a median of \$6.3 K and a mean of \$5.6 K per year. In the study by Piketty, Saez, and Zucman, the mean is even lower at \$1.3 K. Like them, we find a number of zero or even negative values in the ACS data. The zeros are often for young adults who report receiving private transfers, which are not part of the standard ACS income measure that we use. When dropping the negative values or all values below the 1st percentile, the results regarding the incomegradient of public education spending are essentially unchanged.

For post-tax disposable income, we find a mean of \$32.7 K and a median of \$26.3 K based on the ACS data and our simulation using TAXSIM. Piketty, Saez, and Zucman have a mean of \$46.5 K. The difference again arises in the upper half of the income distribution, especially in the top decile. For the bottom 50%, where imputed rents and undistributed profits play not much of a role, our ACS + TAXSIM measure is fairly close to what Piketty, Saez, and Zucman find. With the exception of the bottom decile, our numbers are a bit lower than theirs even for this group, however, despite our use of a more recent year (2017 vs. 2014).

3.3 Public Expenditure on Education

Per-student values Annual public expenditure on education in the United States in 2017 is taken from the OECD's "Education at a Glance" database (OECD Statistics 2020), subsection "Educational finance indicators". The information is available for different levels of education, based on the ISCED 2011 classification (Table 2). Total public expenditure in 2017 is \$56 B at the pre-primary level (ISCED 0), \$296 B at the primary level (ISCED 1), \$328 B at the secondary level (ISCED 2–3), and \$308 B at the tertiary

Table 1: Summary Statistics

		Decile									
	1	2	3	4	5	6	7	8	9	10	Total
A. Pretax income											
Annual Income (Median)	6.3	13.1	18.5	23.8	29.3	35.5	43.2	53.5	70.0	118.0	32.2
Annual Income (Mean)	5.6	13.1	18.5	23.7	29.3	35.6	43.4	53.7	71.0	147.2	43.9
Annual Income (Mean, PSZ)	1.3	9.6	16.0	23.0	31.1	41.3	53.6	69.6	96.8	303.9	64.6
Household Size	2.70	3.10	3.19	3.20	3.12	3.08	2.97	2.86	2.79	2.71	2.97
Children in HH $(0/1)$	0.30	0.35	0.37	0.37	0.36	0.37	0.37	0.38	0.39	0.40	0.37
Age	46.8	50.3	49.3	48.7	48.1	47.7	47.3	47.4	48.2	50.1	48.4
Age Youngest Child	7.4	7.4	7.6	7.9	8.2	8.3	8.3	8.5	8.7	8.8	8.1
B. Posttax cash income											
Annual Income (Median)	6.3	12.5	16.8	20.6	24.6	28.9	34.0	40.3	50.2	78.6	26.7
Annual Income (Mean)	5.6	12.5	16.8	20.7	24.6	28.9	34.0	40.4	50.7	95.8	33.0
Annual Income (Mean, PSZ)	3.5	11.9	17.4	22.4	27.5	33.5	40.8	50.5	66.8	190.3	46.5
Household Size	2.63	3.08	3.21	3.24	3.14	3.07	2.96	2.89	2.79	2.70	2.97
Children in HH $(0/1)$	0.25	0.33	0.38	0.39	0.38	0.38	0.37	0.39	0.39	0.39	0.37
Age	46.7	49.1	47.7	47.4	47.8	47.9	48.1	48.3	49.5	51.4	48.4
Age Youngest Child	7.8	7.6	7.5	7.6	8.0	8.2	8.3	8.4	8.6	8.9	8.1

Notes: The table shows summary statistics for our estimation sample with respect to income and household composition. The income measures are compared with the values reported by Piketty, Saez, and Zucman (2018). The upper panel divides adult (aged 20 and above) into deciles based on their pre-tax income. In the lower panel, the deciles are based on post-tax cash income instead. Pre-tax income is taken directly from the ACS, while post-tax cash income is simulated using TAXSIM. Source: Own calculations based on the American Community Survey 2017. For comparison with the DINA approach, in households with more than one adult, household income is divided by the number of adults (equal-split). The income reported in the table is annual income in thousand US Dollars. N=2,375,184 adult individuals (aged 20 and above). For the sake of presentation and given the large sample size, standard errors are omitted. HH: Household. PSZ (Piketty, Saez, and Zucman 2018): Pre-tax income from Appendix Tables II-B4 (deciles 1-9, computed as the average over percentiles 0-9, 10-19 etc.) and II-B3 (overall mean and decile 10). Post-tax cash income from Appendix Tables II-C4b (deciles 1-9, computed as the average over percentiles 0-9, 10-19 etc.) and II-C3e (overall mean and decile 10). Piketty, Saez, and Zucman report averages for post-tax cash income by percentile of post-tax income (including non-cash transfers), while the averages for post-tax cash income in our ACS data are computed for deciles of post-tax cash income only, because the focus of our paper is to probe the assumption that the non-cash componenents (in our case, education) are distributed proportionally to post-tax cash income. However, as Piketty, Saez, and Zucman allocate non-cash transfers proportionally to post-tax cash income, the deciles for posttax cash income and post-tax income should coincide in their case.

level.¹⁰ In line with the practice of national accounting and the DINA approach, expenditure is valued at cost, as opposed to the valuation that students or their parents put on this expenditure, which is much more difficult to measure.

The numbers are for "all expenditure types" in the OECD nomenclature. This includes both current expenditure (a large share of which are salaries and wages) and capital outlays, but excludes R&D as well expenditure for ancillary services. R&D expenditure is relevant only at the tertiary level, where it amounts to \$37 B in 2017. As part of our robustness checks, we use both a narrower (only current expenditure) and a broader (all expenditure types plus R&D and ancillary services) definition of public education spending. This has little effect on the results, which are mostly driven by differences in enrollment across the income distribution.

The OECD calculates expenditure per student on the basis of full-time equivalents. In these calculations, students in part-time education—relevant only at the pre-primary and the tertiary level—are assumed to represent one-third of a full-time equivalent. Since we do not observe part-time student status in the ACS, we assign the expenditure per full-time equivalent to all students.

Public per-student expenditure in 2017 is around $$13\,\mathrm{K}$ at both the pre-primary and the primary level and slightly higher ($$14.5\,\mathrm{K}$) at the secondary level.¹¹

At the tertiary level, expenditure per student is \$29.1 K. This is an average over 2-Year and 4-Year colleges. As part of our robustness checks, we try to distinguish between the two categories. While the annual per-student expenditure can be calculated by going back to the NCES data on enrollment and expenditure, which is more detailed than what the OECD provides, there is no information in the ACS on the type of college. However, the ACS distinguishes between undergraduate studies on the one hand and graduate and professional schools on the other. In a robustness check, we assign all graduate students to 4-Year colleges, and randomly assign undergraduates to either 2-Year or 4-Year colleges, based on the relative importance of the two types as reported by the NCES.

The education expenditure includes all levels of government—this is important as in both the US and Germany most public education spending occurs at the state and local levels. The OECD only provides the national average of education spending. As part of our robustness checks, we use averages by state provided by the National Center for Education Statistics.¹²

 $^{^{10}\}mathrm{We}$ abstract from post-secondary non-tertiary education, where annual public expenditure in 2017 is a mere \$1.2 B.

¹¹The distribution of funds among the ISCED levels 0, 1, 2, and 3 are estimated by the OECD. The National Center for Education Statistics (De Brey et al. 2021) reports only a single value for these levels. We run a robustness check in which we discard the small differences and use the NCES number.

¹²State-level information is taken from the National Center for Education Statistics, Digest of Education Statistics 2019 (De Brey et al. 2021). For the pre-primary, primary, and secondary levels, we use the values from Table 236.75: Total and current expenditures per pupil in fall enrollment in public elementary and secondary schools, by function and state or jurisdiction: 2016-17. To compute public per-student expenditure at the tertiary level, we divide total expenditure (Table 334.20: Total expenditures of public

Table 2: Public Expenditure on Education

	Per student (\$ K)	Total (\$B)						
	OECD	ACS	OECD	NCES	NIPA			
Pre-primary	13.0	38	56	n.a.	n.a.			
Primary	13.0	283	296	n.a.	n.a.			
Secondary	14.5	384	328	n.a.	n.a.			
Sum		705	681	681	666			
Tertiary	29.1	488	308	335	288			
Total		1.192	990	1.016	954			

Notes: The table reports the per-student values for annual public expenditure on education that we use in our calculations (column 1) and the aggregates that we find when combining these values with our ACS enrollment data (column 2). These aggregates are compared with statistics published by the OECD, the National Center for Education Statistics (NCES), and National Accounts (NIPA). Sources: OECD: OECD Statistics (2020). Postsecondary non-tertiary education (ISCED level 4) is negligible and omitted for simplicity. OECD per-student expenditure is for full-time equivalents. The OECD assumes that part-time students receive one-third of a full-time equivalent. The distinction between full-time and part-time is only relevant at the pre-primary and the tertiary levels. ACS: Own calculations based on enrollment as observed in the American Community Survey 2017, combined with the per-student expenditure numbers of the OECD. NCES: National Center for Education Statistics, Digest of Education Statistics 2019 (De Brey et al. 2021), Table 236.10: Summary of expenditures for public elementary and secondary education and other related programs, by purpose: Selected years, 1919-20 through 2016-17. NIPA: Table 3.16. Government Current Expenditures by Function (Data published on March-26-2021 in connection with the Third estimates for 2020 Q4). NIPA Codes: Total: G16029; Elementary and secondary: G16030; Tertiary = Higher (G16031; $195 \, B + Libraries (G16032; 13 \, B) and Other (G16033; 81 \, B).$

Remarks: (1) The US national sources only report aggregates for elementary and secondary education. The breakdown into pre-primary, primary, and secondary is estimated by the OECD. (2) Piketty, Saez, and Zucman (2018) report a lower value for education spending, which corresponds to current expenditure only (NIPA code G17019). The value reported in their paper is \$762 B in 2014 (see Section 2 above). The NIPA value for 2014 has since been updated to \$789 B. In the 2020 update of their analysis, Piketty, Saez, and Zucman use the same measure that we employ in this paper (NIPA Code G16029) and report values of \$884 B for 2014 and \$956 B for 2017 (see PSZ2020AppendixTablesI(Aggreg).xlsx, Sheet DataIncome, Column PF, available at http://gabriel-zucman.eu/usdina/, which is almost identical to the \$954 B reported in the table, the small difference being likely due to an update of the NIPA data.

Unfortunately, we do not have data on per-capita public expenditure at the sub-state level that would allow us to capture differences between richer and poorer school districts or neighborhoods. This means that the differences in public spending by income that we document are driven by different enrollment rates, different propensities to choose public vs. private institutions, and, in the robustness check, by differences across states. We do not capture any remaining income-related variation in per-capita spending. As this remaining variation is likely positively related to income, this means that we do not capture one component that would work towards the proportionality assumption used as the benchmark in the DINA approach. However, we find such a strong departure from proportionality that the within-state differences in per-capita spending would have to be implausibly high in order to justify the assumption. Moreover, at least at the level of school districts, the difference by income is less pronounced than one might think, and is characterized by a U-shape instead of a monotonous increase with income. Average perpupil expenditure in public elementary and secondary schools is \$12.9 K in low-poverty districts, \$11.2 K in middle-low poverty districts, \$10.8 K in middle-high poverty districts, and \$13.0 K in high-poverty districts (De Brey et al. 2021, Table 236.85). There is a ruralurban divide: while in cities high-poverty districts have substantially higher public perpupil spending than low-poverty districts, the difference is smaller in suburban districts and turns in favor of low-poverty districts in towns and rural areas.

Aggregates Table 2 also shows aggregate annual expenditure. Our own numbers—obtained from combining the enrollment observation in the ACS with the OECD values for per-student expenditure—are compared with the OECD aggregates, information from the National Center for Education Statistics (NCES), and with the national accounts (NIPA) data published by the Bureau of Economic Analysis, which is the source that Piketty, Saez, and Zucman (2018) use. (They only report the total, without the breakdown by education level.)

At ISCED levels 0-3 (pre-primary, primary, secondary), we obtain an aggregate expenditure of \$705 B, close to the \$681 B reported by both the OECD and the NCES, and only about 5% higher than the NIPA figure of \$666 B. That our value is slightly higher than the OECD figure is due to two factors. First, as shown in Figure 13, the ACS enrollment numbers are slightly higher than what is reported by the OECD (51.4 M vs. 50.6 M). Second, some of the children in pre-primary education attend kindergarten only part of the day. When computing full-time equivalents, the OECD assigns individuals in part-time education a weight of 0.3. In the ACS, we do not observe part-time status, and assign all individuals the full-time equivalent expenditure reported by the OECD. This

degree-granting postsecondary institutions, by level of institution, purpose of expenditure, and state or jurisdiction: 2014-15 through 2017-18) by the number of students (Table 304.15: Total fall enrollment in public degree-granting postsecondary institutions, by state or jurisdiction: Selected years, 1970 through 2018).

amounts to the assumption that all individuals are in fact in full-time education, which leads us to overestimate the annual expenditure.

Both factors are aggravated at the tertiary level. In the ACS, there are 16.8 M students enrolled in public tertiary institutions, while the OECD and the NCES report 14.6 M students, a difference of 2.2 M or about 15% (Figure 13). Moreover, the part-time share is even higher than for pre-primary education. As a result of both factors, our estimate of annual public expenditure at the tertiary level of \$488 B is substantially higher than the numbers reported by the OECD, the NCES, and the Bureau of Economic Analysis, which range between \$288 B and \$335 B.

As part of our robustness checks, we address these issues by rescaling the enrollment numbers in the ACS so that we have the same number of full-time equivalent students as the OECD. We do this both in a neutral way—by assuming that the excess number of full-time equivalents is independent of income—and as a bounds analysis in which we assume that the excess mass is concentrated in either the bottom or the top half of the income distribution.

4 Results

4.1 Distribution of Public Education Spending

Allocation Based on Actual Enrollment Figure 6 shows how public education spending is the United States in 2017 is distributed among the deciles of the income distribution. Following Piketty, Saez, and Zucman (2018), the distribution is for adults age 20 and above; in households with more than one adult, income is split equally. Income is pre-tax income as reported in the American Community Survey; below, we report results when we use simulated post-tax income instead, as a first step towards the more comprehensive measure of post-tax cash income used by Piketty, Saez, and Zucman.

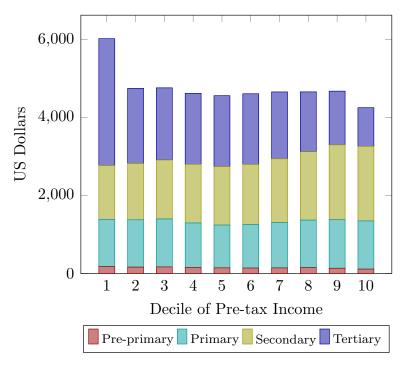
Public education spending is highest in the first decile—with an average of \$6.0 K per adult—and lowest in the tenth decile of the pre-tax income distribution, where the average is \$4.3 K. In deciles 2 to 9, the means of per-capita spending are fairly close together, at between \$4.5 K and \$4.8 K. The overall average is \$4.8 K. The Bottom 50% of the pre-tax income distribution receive an average of \$4.9 K per year in terms of public

¹³Part of the difference is probably due to students at private non-profit institutions. According to the NCES, 1.1 M students attended such an institution in 2017. If some of these declared to be in a public institution in the ACS because they equated not-for-profit with public, this could explain part of the higher number of students at public institutions that we find. Note, however, that we also overestimate the total number of students at the tertiary level, so the measurement issue does not only concern the classification of institutions into public or private.

 $^{^{14}}$ According to the OECD, 1.6 M out of 5.1 M children (68%) in pre-primary education attend kinder-garten only part-time. At the tertiary level, there are 6.4 M part-time students (43% of the total 14.6 M). At the primary and secondary levels, all pupils attend school full-time.

¹⁵The numerical values are reported in Table 4 in the Appendix.

Figure 6: Public Education Spending by Pre-tax Income, Allocated Based on Actual Enrollment



Notes: The figure shows how public education spending in the United States in 2017 is distributed among the deciles of the pre-tax income distribution. For each decile, the bars show the average values of annual public education spending (in 2017 US Dollars) at the pre-primary, primary, secondary, and tertiary levels of education. Source: Enrollment in public educational institutions is taken from the American Community Survey 2017. Each pupil or student is assigned the per-capita value of public education spending taken from the OECD (see Table 2). Public education expenditure is summed up at the household level, and the resulting sum is split equally among adults aged 20 and above in the household. Household income is likewise split equally among all adults.

education spending, followed by the Middle 40% with $$4.7\,\mathrm{K}$, and, as noted, the Top 10% with $$4.3\,\mathrm{K}$.

Turning to the different levels of education, we see little differences by income for preprimary and primary education. Per-capita expenditure on secondary education tends to grow with income, with an average of \$1.4 K allocated to each adult in decile 1 and about \$1.9 K in deciles 9 and 10. Public spending on tertiary education shows the opposite pattern. It is the driver behind the regressivity of public education spending, being concentrated in the bottom decile of the income distribution, where average annual spending is \$3.3 K, more than three times the average in the top decile (\$1.0 K). The high average in the poorest decile is mostly explained by college students who no longer live with their parents. By contrast, the public expenditure on students who are still in the parental household is spread out much more evenly across the income distribution.

Comparison with Proportional and Lump-Sum Allocations Figure 7 contrasts the actual distribution based on the American Community Survey with the proportional allocation used by Piketty, Saez, and Zucman (2018). For public spending on education

(about 30% of collective expenditure in the United States), a proportional allocation is clearly not a good assumption. It implies annual per-capita spending of \$0.6 K in the poorest decile, only a tenth of the value that we find based on actual enrollment data from the American Community Survey. At the top of the income distribution, the proportionality assumption allocates \$18.4 K to each adult in the richest decile, more than four times the value based on the ACS. Furthermore, as pointed out in Section 2, given the unequal distribution of pre-tax income even within the top decile, a proportional allocation implies implausibly high per-capita values among individuals in, say, the Top 1% or Top 0.1% of the distribution.

As microdata on enrollment in education is easily available for the United States and other countries, we believe that the precision of the DINA approach can be improved at little cost by replacing the proportionality assumption with an allocation based on actual enrollment. An even easier fix consists in replacing the allocation proportional to post-tax cash income—which the DINA Guidelines recommends as the benchmark—by a lump-sum allocation. As Figure 7 shows, assigning the mean of \$4,783 to each adult is a good approximation to the distribution based on actual enrollment.

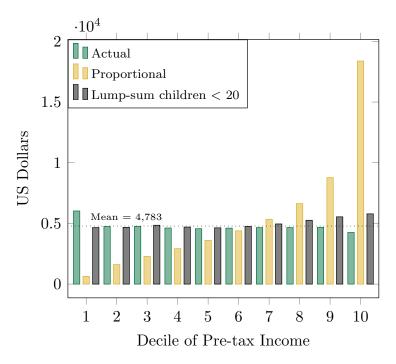
Figure 7 also includes the distribution that arises from allocating public education spending as a lump-sum transfer per child below the age of 20, as in the robustness check in the paper by Piketty, Saez, and Zucman (2018). This assumption performs much better than the proportional allocation. The differences with respect to our baseline results arise from the fact that this shortcut method does not take into account the differences in percapita expenditure by level of education (tertiary education is much more expensive than the rest, at least in the United States), and especially that it does not capture public spending that goes to college students age 20 and above.

Regressivity Driven by Age Effects The regressivity of public education spending in the cross-section is strongly driven by age effects (Figure 8). Individuals aged 20 to 24 receive a lot of education spending on average, mostly for their own (tertiary) education. At the same time, they have by far the lowest current income of all age groups. Preprimary and primary education does not play a large role in this age group, as the share of parents is still low. Spending on secondary education is a bit higher because some individuals are still in secondary education themselves.

In the age group 25–29, average public education spending is much lower. (Own) tertiary education is still significant, but less so than for individuals in their early 20s. Secondary education also drops in importance, while public expenditure on pre-primary and primary education starts building up as individuals in this group have more (and older) children than in the age group just below.

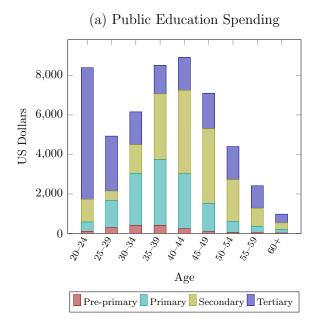
In the older age groups, the share of parents and the age of their children continue to rise, as reflected in the increasing public expenditure at the pre-primary, primary, and

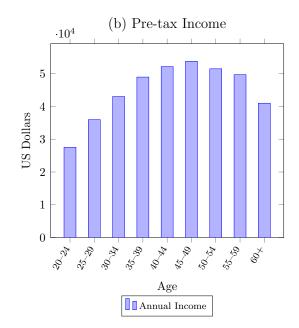
Figure 7: Public Education Spending by Pre-tax Income: Comparison of Allocation Methods



Notes: The figure compares the actual distribution of public education spending (in green, this is the same distribution as in Figure 6) with the distributions that result from an allocation that is proportional to pre-tax income (yellow) and from a lump-sum transfer (gray) to all children below age 20, irrespective of actual enrollment and disregarding the differences in per-capita spending between pre-primary, primary, secondary, and tertiary education. The figure also shows the value of \$4,783 that would result from a lump-sum allocation to all adults. Source: Enrollment in public educational institutions is taken from the American Community Survey 2017. Each pupil or student is assigned the per-capita value of public education spending taken from the OECD (see Table 2). Public education expenditure is summed up at the household level, and the resulting sum is split equally among adults aged 20 and above in the household. Pre-tax household income is likewise split equally among all adults.

Figure 8: Public Education Spending and Pre-tax Income by Age





Notes: The left panel of the figure shows how the average value of public education spending (allocated based on actual enrollment) differs by age. The right panel depicts average pre-tax income for the same age categories. Source: Enrollment in public educational institutions is taken from the American Community Survey 2017. Each pupil or student is assigned the per-capita value of public education spending taken from the OECD (see Table 2). Public education expenditure is summed up at the household level, and the resulting sum is split equally among adults aged 20 and above in the household. Pre-tax household income is likewise split equally among all adults.

secondary levels. While the first two peak in the age group 35–39, spending on secondary and tertiary education continues into age groups 40-44 and 45–49, respectively. At later ages, expenditure falls for them as well as children leave the parental household. The maximum of total public education spending is reached in the age group 40–44. Pre-tax income, by contrast, peaks at age 45–49, and is still fairly high thereafter while public education spending declines steeply for individuals in their late 40s and in their 50s. Together with the high level of public education spending for the poorest age group 20–24, this phase shift drives the regressivity of public education spending in the cross-section.

4.2 Robustness Checks

Post-tax Cash Income So far, our results have been for pre-tax income, which is directly observable in the American Community Survey. However, Piketty, Saez, and Zucman (2018) assume that education and other collective expenditure items are allocated proportionally to post-tax cash income. We therefore run a robustness check in which we use our measure of post-tax cash income—simulated using TAXSIM—to divide adults into deciles (Figure 14 and Table 4 in the Appendix). As for pre-tax income, the proportionality assumption is rejected, while a lump-sum allocation is a reasonable

approximation except for the bottom and the top of the income distribution.

Household Equivalence Income As noted in Section 1, there is an older literature that augments the standard survey measures of disposable (money) income by different components of public in-kind spending, often with a cross-country focus. These studies measure income at the household level and attempt to make households of different size and age composition comparable through equivalence scales. As Figure 15 in the Appendix shows, public education spending remains regressive when adopting such a household perspective. The average amount of public education spending received is now higher as the transfers are measured at the household level and not divided equally among adults. When the deciles are defined based on pre-tax income, average spending declines throughout the distribution. For a distribution based on post-tax cash income, a lump-sum allocation is a decent approximation for the bottom three or four deciles, but the upper half of the distribution is again characterized by a negative relationship between public education spending and household income.

The finding that public education spending declines with household income is in line with the study by Zwijnenburg, Bournot, and Giovannelli (2017) who report the percentage of total education spending by quintiles of household disposable income for the United States and several other countries. In the United States in 2012, 25.4% of public education spending goes to households in the bottom quintile, compared with 15.3% in the top quintile. In our data for 2017, the shares are similar, but the regressivity is even more pronounced: 26.3% of public spending goes to the 20% of households with the lowest post-tax cash income, while the richest 20% receive 11.1% of the total.

Other Checks We also ran a number of other, more technical robustness checks. As noted above, despite the legal obligation to answer the survey, some of the individual income components are actually imputed by the data provider. When we drop all households in which more than half of household income is based on an imputation (slightly less than 20% in our sample), the results are virtually unchanged (Table 5). The same holds when we drop all households with negative income or all households with income below the 1st percentile. When the threshold is increased to the 2.5th percentile, average public education spending in the first decile is reduced from \$6.0 K to \$5.4 K, but is still higher than in all other deciles. Dropping all households whose income is above the 99.5th percentile likewise has no effect on the results.

As seen in Figure 13, the ACS slightly overestimates the enrollment in educational institutions by comparison with the numbers reported by the OECD and the NCES. When we scale down our ACS enrollment numbers to meet the NCES numbers, average

¹⁶The numerical values are reported in Table 4 in the Appendix.

public education spending goes down in all deciles, but the negative relationship with income is preserved.

In our main specification, we use a single value for per student expenditure at the different levels of education, as the OECD does not provide information on within-country variation. When we use state-specific values from the NCES instead, the difference between the first and the tenth deciles is slightly reduced, but the poorest decile still receives substantially more public education spending. A lump-sum allocation is again a good approximation for the deciles in between.

The OECD calculates expenditure per student on the basis of full-time equivalents; students in part-time education—relevant only at the pre-primary and the tertiary level—are assumed to represent one-third of a full-time equivalent. In the ACS, there is no information on whether individuals are enrolled only part-time, and we assign the expenditure per full-time equivalent to all students in our main specification. As a robustness check, we randomly assign part-time status based on the share of part-time students reported by the OECD. This brings down the average expenditure by decile, but leaves the negative income gradient intact.

The OECD only reports a single number for annual per-student expenditure at the tertiary level, which is an average over 2-Year and 4-Year colleges. In a robustness check, we assign all graduate students to 4-Year colleges, and randomly assign undergraduates to either 2-Year or 4-Year colleges, based on the relative importance of the two types as reported by the NCES. Average expenditure is higher than in the main specification, but the relationship between income and expenditure remains the same.

In our main specification, public expenditure per student includes both current expenditure (a large share of which are salaries and wages) and capital outlays, but excludes R&D—which is relevant only at the tertiary level—as well expenditure for ancillary services. Alternatively, we have used a narrower (only current expenditure) and a broader (all expenditure types plus R&D and ancillary services) definition of public education spending. This changes the level of expenditure, but has little impact on the income gradient.

4.3 Beyond the Cross-Section

The DINA literature invokes two arguments for assigning public education spending proportionally to post-tax cash income: the unequal access to education by parental income—e.g., Alvaredo et al. (2020, p. 65) or Saez and Zucman (2020, p. 33)—and "a lifetime perspective where everybody benefits from education, and where higher earners attend better schools and for longer" (Piketty, Saez, and Zucman 2017, p. 27/28).

In the following, we show that individuals with higher lifetime earnings (proxied for by earnings at age 40–45) have indeed received substantially more public education spending

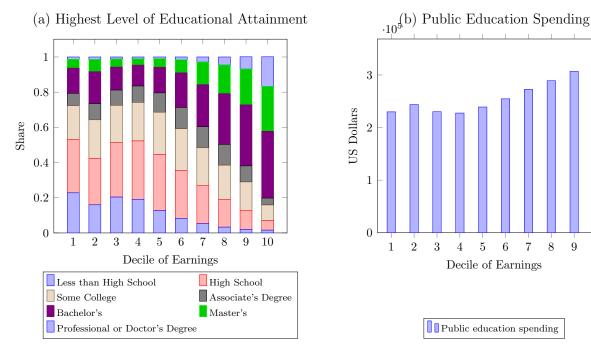
in the past. We also show—based on data from Germany—that more public education spending goes to children whose parents have a higher socio-economic status. In both cases, we depart from the cross-sectional perspective we have adopted so far. In particular, we do not consider the public education spending received in a single year, but the sum of spending received in the education system. We classify individuals by their highest degree and assume that a given degree implies that the individual has passed through all the stages below, and that everyone needed the same number of years to complete each stage. 17 This is admittedly a simplification. For instance, not every child attends kindergarten, and some students repeat a year in school or take longer to finish a bachelor's or master's degree, and this variation is likely correlated with both lifetime earnings and parents' socio-economic status. However, with our data there is little we can do about this, and the differences that we find are so large that they are robust to different assumptions. A potentially more important qualification is that we do not observe whether individuals completed their education abroad. We have no information about this in our data, and assume that the entire schooling was obtained in the United States. Another shortcut that we take is to use the 2017 per-student values for public education expenditure (Table 2) although the cohort of individuals that we consider—40-45-year-olds in 2017, i.e. people born in the early and mid-1970s—obtained their education in the past. Given that we consider a cohort of only six years and that our interest is in the gradient and not the level of spending, this assumption should be fairly innocuous as well. Finally, moving beyond the cross-section—i.e., current educational enrollment—means that we cannot distinguish between public and private institutions anymore. We assume that all individuals obtained their degrees in the public education system. This means that we overestimate the level of expenditure and, more importantly, the income gradient, as graduating from a private college is positively correlated with both own lifetime earnings and parents' socio-economic status.

Differences by Lifetime Earnings We proxy for lifetime earnings using current earnings of individuals aged 40–45. At this age, the rank correlation between current earnings and lifetime earnings reaches its maximum (e.g. Bönke, Corneo, and Lüthen 2015; Haider and Solon 2006). As we now consider earnings and not income, we do not use the equal-split assumption that we adopt in the cross section, but directly use the personal earnings information available in the ACS.

Figure 9 shows how the highest degree and public education spending received vary with earnings. As expected, the highest degree is positively correlated with earnings (Panel a). While in the bottom half of the earnings distribution most individuals have at most a high school diploma or attended college without obtaining a degree, the share

 $^{^{17}}$ The details of our mapping between the highest degree observed in the ACS and the number of years spent at the different ISCED levels are presented in Table 6 in the Appendix.

Figure 9: Highest Degree and Public Education Spending by Current Earnings, Individuals Aged 40-45



Notes: The figure shows the highest degree (left panel) and public education spending by current earnings (right panel) for individuals aged 40–45. Source: Own calculations based on the American Community Survey 2017. When calculating public education spending, we assume that a given degree implies that the individual has passed through all the stages below, and that everyone needed the same number of years to complete each stage (see Table 6 in the Appendix for details). We also assume that all individuals have attended only public educational institutions. Each year in the education system is multiplied with the per-capita value of public education spending taken from the OECD (see Table 2). We use the 2017 values of per-capita spending although the individuals who were 40-45 years old in 2017 obtained their education in earlier years.

9 10

of people with a bachelor's, master's, or professional and doctor's degree increases in the upper half of the earnings distribution.

When translating these differences in degrees into differences in public education spending received, there is—unlike in the cross-section—a positive income (or, more precisely, earnings) gradient. The 10% of individuals with the highest earnings have received average public education spending of \$335 K, about 1.4 times the amount of the bottom 50% (\$234 K). The allocation is still not proportional to earnings, however; proportionality would require a factor of about 14 (\$196 K vs. \$14 K).

Intergenerational Perspective The second argument invoked in the DINA literature for assigning public education spending proportionally to post-tax cash income is the unequal access to education by parents' income or, more generally, socio-economic status (SES). Studies documenting this inequality are legion. Children from a more advantaged socio-economic background tend to go to better schools and are more likely to attend college. We show that these differences indeed produce a positive relationship between

parents' SES and the public education expenditure that their children receive. Unfortunately, we do not observe parents' SES in the American Community Survey. We therefore make use of the German Socio-Economic Panel (SOEP) instead, which we again combine with information from the OECD on current public expenditure per student.¹⁸

Like in the analysis based on lifetime earnings, we restrict the sample to individuals aged 40–45 in 2017. As we do not observe individual trajectories, we again assume that individuals followed a stylized path to their highest degree (no grade retentions etc.). However, we do take into account differences between East and West Germany in kindergarten enrollment, which was and still is higher in the East.

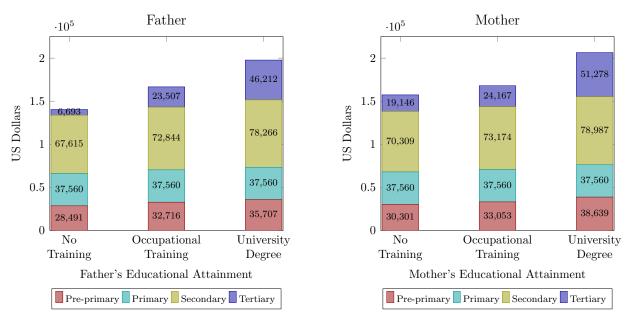
Figure 10 shows that individuals whose parents had a university degree received substantially more public education spending than children of parents with occupational training or no training at all. These differences arise at all levels except primary school, which lasts four years for all children. The differences at the pre-primary level are driven exclusively by the East-West difference in enrollment. Given that for the cohort considered here (born between 1972 and 1977), we construct our enrollment data using interpolations based on sparse historical data, these differences should be interpreted cautiously. More important are the differences at the secondary level, which ends at grade 9 for individuals in the general track and, depending on the federal state, at grade 12 or 13 for people who finished the academic track. The differences are even larger at the tertiary level. Individuals whose father had no occupational training hardly attended college, resulting in a very low (unconditional) average of public education spending at the tertiary level (\$6.7 K). By contrast, individuals whose father had a university degree received a much higher \$46.2 K on average. Among individuals whose mother had a university degree is even higher, at \$51.3 K.

Total public spending on education is on average \$170 K for the cohort considered here. The difference between individuals from the most and the least privileged background with respect to parents' education is \$57 K on the father's side and \$49 K on the mother's side. For parents' occupational prestige (measured using the occupation that they held when the respondents were 15 years old), the differences are slightly less pronounced, but still sizeable. Individuals whose father's occupation is in the top quartile of the Standard International Occupational Prestige Score (SIOPS) received, on average, \$36 K more public education spending than individuals with fathers in the bottom quartile; on the mother's side, the difference is \$33 K (Figure 16 in the Appendix).

DINA as a Cross-Sectional Approach That children from already more privileged backgrounds receive \$50 K more in public education expenditure is more important for the distributional debate than the regressive pattern of public expenditure found in any given year, which, as seen above, is strongly driven by age effects. However, the positive

¹⁸The data are described in Section A in the Appendix.

Figure 10: Public Education Spending by Parents' Education. Germany, Individuals Aged 40–45



Notes: The figure depicts average public education spending by parents' education for individuals aged 40–45 in Germany. The left panel distinguishes by the education of the father, the right panel by the education of the mother. Source: Own calculations using the Socio-economic Panel (SOEP) 2017. Public expenditure per student is taken from the "Education at a Glance" database (OECD Statistics 2020). Euro amounts were converted into US Dollars using the GDP purchasing power parities provided by the OECD.

association between public expenditure and parental SES or own lifetime earnings does not provide a justification for allocating public education expenditure proportionally to income in the DINA approach. So far, the approach has been exclusively cross-sectional, and departing from this cross-sectional perspective only for public education spending seems ad hoc. After all, age effects are also present in earnings or capital income, but are not adjusted for when measuring pre-tax income. Likewise, many cash transfers such as family benefits or in-kind transfers such as Medicare are also age-dependent, but are assigned to current recipients in the DINA approach.

5 Conclusion

In the distributional national accounts (Dina) created by Piketty, Saez, and Zucman (2018) and others, government collective expenditure (e.g., education, defense, infrastructure) is typically allocated proportionally to post-tax cash income, which renders half of government expenditure distributionally neutral and implies large differences in the per-capita value of collective expenditure. We revisit the study by Piketty, Saez, and Zucman and show that the level of post-tax inequality is fairly sensitive to this assumption. When the expenditure is allocated on a lump-sum basis instead—an assumption

that the recent version of the DINA Guidelines (Alvaredo et al. 2020) suggests as an alternative to the proportional allocation—the gap in post-tax income shares between the Top 10% and Bottom 50% is reduced by half. The trend in US post-tax income shares is hardly affected by the assumptions, however.

The main contribution of our paper is to provide evidence on how an important part of collective expenditure is actually distributed. We find that, when adopting the cross-sectional perspective of the DINA approach, public education spending goes disproportionately to the bottom half of the income distribution. This pattern is strongly driven by age effects. There is indeed a positive relationship between public education spending and lifetime earnings or parents' socio-economic status, but even the relationship with earnings is far from being proportional. More importantly, the last two patterns do not provide an empirical basis for the cross-sectional DINA approach. Adjusting for age effects only for public education, but not for other items such as earnings, capital income, family cash transfers, or Medicare, would introduce an inconsistency into the framework.

Based on our findings, we conclude that public education expenditure should not be allocated proportionally to post-tax cash income as recommended in the DINA Guidelines. As microdata on education is widely available, an allocation based on actual enrollment can improve the distributional analysis of post-tax income at little extra cost. This recommendation is in line with the OECD–Eurostat Expert Group on Disparities in a National Accounts framework (EG DNA), which also argues for an allocation based on actual use (Zwijnenburg 2019). An even easier improvement is to allocate public education spending as a lump-sum transfer, which—at least in the US context of 2017—provides a good approximation of the actual distribution.

Given that a proportional allocation implies very high per-capita values for individuals with high incomes, we believe that a lump-sum allocation is the better benchmark for the remaining parts of government consumption expenditure (defense, infrastructure) as well. However, given the difficulty of assigning these other items to households and individuals, reporting results for both a lump-sum and a proportional allocation is probably a reasonable compromise. Another option is to resort to an income concept such as disposable personal income that takes only money transfers and certain in-kind transfers such as Medicare and Medicaid into account while avoiding the assignment of government consumption expenditure altogether (Gindelsky 2021). Whether this or the more comprehensive DINA income concept is more useful depends on the question at hand.

In our analysis, differences in public education spending result from differences in enrollment and in the choice of public or private institutions. In a robustness check, we also exploit differences in average spending across states. We do *not* capture any remaining income-related variation in per-capita spending. As this remaining variation is likely positively related to income, this means that we do not capture one component that would work toward the proportionality assumption used as the benchmark in the DINA

approach. However, we find such a strong departure from proportionality that the withinstate differences in per-capita spending would have to be implausibly high in order to justify the assumption. Moreover, at least at the level of school districts, the difference by income is less pronounced than one might think, and is characterized by a U-shape instead of a monotonous increase with income. Still, incorporating more fine-grained information on per-capita spending would further increase the precision of the DINA approach and is a useful direction for future research.

References

- Aaberge, Rolf, Manudeep Bhuller, Audun Langørgen, and Magne Mogstad (2010). "The Distributional Impact of Public Services When Needs Differ". In: *Journal of Public Economics* 94, pp. 549–562.
- Aaberge, Rolf, Lasse Eika, Audun Langørgen, and Magne Mogstad (2019). "Local Governments, In-kind Transfers, and Economic Inequality". In: *Journal of Public Economics* 180, p. 103996.
- Aaberge, Rolf, Audun Langørgen, and Petter Lindgren (2013). "The Distributional Impact of Public Services in European Countries". Statistics Norway Research Department Discussion Paper No. 746.
- Alvaredo, Facundo, Anthony B. Atkinson, Thomas Blanchet, Lucas Chancel, Luis Bauluz, Matthew Fisher-Post, Ignacio Flores, Bertrand Garbinti, Jonathan Goupille-Lebret, Clara Martínez-Toledano, Marc Morgan, Theresa Neef, Thomas Piketty, Anne-Sophie Robilliard, Emmanuel Saez, Li Yang, and Gabriel Zucman (2020). "Distributional National Accounts Guidelines: Methods and Concepts Used in the World Inequality Database". Paris. World Inequality Lab.
- Auten, Gerald and David Splinter (2019). "Income Inequality in the United States: Using Tax Data to Measure Long-term Trends". Working Paper.
- Bach, Stefan, Charlotte Bartels, and Theresa Neef (2021). "Distributional National Accounts: A Macro-Micro Approach to Inequality in Germany". DIW and Freie Universität Berlin.
- Blanchet, Thomas, Lucas Chancel, and Amory Gethin (2019). "How Unequal is Europe? Evidence from Distributional National Accounts, 1980–2017". WID.world Working Paper No. 2019/06.
- Bönke, Timm, Giacomo Corneo, and Holger Lüthen (2015). "Lifetime Earnings Inequality in Germany". In: *Journal of Labor Economics* 33 (1), pp. 171–208.
- Bozio, Antoine, Bertrand Garbinti, Jonathan Goupille-Lebret, Malka Guillot, and Thomas Piketty (2018). "Inequality and Redistribution in France, 1990–2018: Evidence from Post-Tax Distributional National Accounts (DINA)". WID.world Working Paper No. 2018/10.
- Callan, Tim, Timothy Smeeding, and Panos Tsakloglou (2008). "Short-run Distributional Effects of Public Education Transfers to Tertiary Education Students in Seven European Countries". In: *Education Economics* 16 (3), pp. 275–288.
- De Brey, Cristobal, Thomas D. Snyder, Anlan Zhang, and Sally A. Dillow (2021). "Digest of Education Statistics 2019 (NCES 2021-009)". Washington, DC. National Center for Education Statistics, Institute of Education Sciences, U.S. Department of Education.
- Feenberg, Daniel and Elisabeth Coutts (1993). "An Introduction to the TAXSIM Model". In: Journal of Policy Analysis and Management 12 (1), pp. 189–194.

- Garbinti, Bertrand, Jonathan Goupille-Lebret, and Thomas Piketty (2018). "Income Inequality in France, 1900–2014: Evidence from Distributional National Accounts". In: *Journal of Public Economics* 162, pp. 63–77.
- Garfinkel, Irwin, Lee Rainwater, and Timothy Smeeding (2006). "A Re-examination of Welfare States and Inequality in Rich Nations: How In-kind Transfers and Indirect Taxes Change the Story". In: *Journal of Policy Analysis and Management* 25 (4), pp. 897–919.
- Gillespie, W. Irwin (1965). "Effect of Public Expenditures on the Distribution of Income". In: *Essays in Fiscal Federalism*. Ed. by Richard A. Musgrave. Washington, DC: Brookings Institution, pp. 122–186.
- Gindelsky, Marina (2021). "Do Transfers Lower Vertical and Horizontal Inequality? Evidence from Distributional National Accounts". Bureau of Economic Analysis, mimeo.
- Goebel, Jan, Markus M. Grabka, Stefan Liebig, Martin Kroh, David Richter, Carsten Schröder, and Jürgen Schupp (2019). "The German Socio-Economic Panel (SOEP)". In: Jahrbücher für Nationalökonomie und Statistik 239 (2), pp. 345–360.
- Haider, Steven and Gary Solon (2006). "Life-Cycle Variation in the Association between Current and Lifetime Earnings". In: *American Economic Review* 96 (4), pp. 1308–1320.
- Hammar, Olle, Paula Roth, Daniel Waldenström, and Gabriel Zucman (2020). "Distributional National Accounts in the Welfare State: Sweden, 1930–2016". In progress.
- Horton, Tim and Howard Reed (2010). Where the Money Goes: How We Benefit from Public Services. Trade Unions Congress.
- Jestl, Stefan and Emanuel List (2020). "Distributional National Accounts (DINA) for Austria, 2004–2016". World Inequality Lab Working Paper, 2020/23.
- Marical, Francois, Marco Mira d'Ercole, Maria Vaalavou, and Gerlinde Verbist (2006). "Publicly-provided Services and the Distribution of Resources". In: *OECD Social, Employment and Migration Working Papers* (45).
- Musgrave, Richard A., Karl E. Case, and Herman Leonard (1974). "The Distribution of Fiscal Burdens and Benefits". In: *Public Finance Quarterly* 2 (3), pp. 259–311.
- O'Dea, Cormac and Ian Preston (2012). "The Distributional Impact of Public Spending in the UK". London. IFS Working Paper W12/06.
- O'Higgins, Michael and Patricia Ruggles (1981). "The Distribution of Public Expenditures and Taxes among Households in the United Kingdom". In: *Review of Income and Wealth* 27 (3), pp. 298–326.
- OECD Statistics (2020). Education at a Glance Dataset "Educational finance indicators". URL: https://stats.oecd.org (visited on 09/23/2021).
- Paulus, Alari, Holly Sutherland, and Panos Tsakloglou (2010). "The Distributional Impact of In-Kind Public Benefits in European Countries". In: *Journal of Policy Analysis and Management* 29 (2), pp. 243–266.

- Piketty, Thomas, Emmanuel Saez, and Gabriel Zucman (2017). Distributional National Accounts: Methods and Estimates for the United States. Data Appendix. URL: http://gabriel-zucman.eu/usdina/ (visited on 09/23/2021).
- (2018). "Distributional National Accounts: Methods and Estimates for the United States". In: *Quarterly Journal of Economics* 133 (2), pp. 553–609.
- Piketty, Thomas, Li Yang, and Gabriel Zucman (2019). "Capital Accumulation, Private Property, and Rising Inequality in China, 1978–2015". In: *American Economic Review* 109 (7), pp. 2469–2496.
- Ruggles, Patricia and Michael O'Higgins (1981). "The Distribution of Public Expenditure among Households in the United States". In: *Review of Income and Wealth* 27 (2), pp. 137–164.
- Ruggles, Steven, Sarah Flood, Ronald Goeken, Josiah Grover, Erin Meyer, Jose Pacas, and Matthew Sobek (2020). *IPUMS USA: Version 10.0 [dataset]*. Minneapolis, MN: IPUMS. URL: https://doi.org/10.18128/D010.V10.0.
- Saez, Emmanuel and Gabriel Zucman (2016). "Wealth Inequality in the United States since 1913: Evidence from Capitalized Income Tax Data". In: *Quarterly Journal of Economics* 131 (2), pp. 519–578.
- (2020). "Trends in US Income and Wealth Inequality: Revising After the Revisionists". NBER Working Paper No. 27921.
- Samwick, Andrew A. (2013). "Donating the Voucher: An Alternative Tax Treatment of Private School Enrollment". In: *Tax Policy and the Economy* 27 (1), pp. 125–160.
- Smeeding, Timothy (1977). "The Antipoverty Effectiveness of In-Kind Transfers". In: *The Journal of Human Resources* 12 (3), pp. 360–378.
- Smeeding, Timothy, Peter Saunders, John Coder, Stephen Jenkins, Johan Fritzell, Aldi J. M. Hagenaars, Richard Hauser, and Michael Wolfson (1993). "Poverty, Inequality, and Family Living Standards Impacts across Seven Nations: The Effect of Noncash Subsidies for Health, Education and Housing". In: *Review of Income and Wealth* 39 (3), pp. 229–256.
- Smith, Matthew, Danny Yagan, Owen Zidar, and Eric Zwick (2019). "Capitalists in the Twenty-first Century". In: *Quarterly Journal of Economics* 134 (4), pp. 1675–1745.
- Verbist, Gerlinde, Michael Förster, and Maria Vaalavou (2012). "The Impact of Publicly Provided Services on the Distribution of Resources: Review of New Results and Methods." In: OECD Social, Employment and Migration Working Papers (130).
- Zwijnenburg, Jorrit (2019). "Unequal Distributions: EG DNA versus DINA Approach". In: AEA Papers and Proceedings 109, pp. 296–301.
- Zwijnenburg, Jorrit, Sophie Bournot, and Federico Giovannelli (2017). "OECD Expert Group on Disparities in a National Accounts Framework Results from the 2015 Exercise". Paris. OECD Statistics Directorate Working Paper No. 76.

Appendix

A Supplementary Analyses: Data for Germany

German Socio-Economic Panel For some of the supplementary analyses, we draw on data for Germany from the Socio-Economic Panel (SOEP). The SOEP is a well-established panel study that began in 1984 (Goebel et al. 2019). As with the ACS, we use the 2017 wave. The SOEP is much smaller than the ACS, but has the advantage that it allows analyzing a respondent's own educational attainement by parental education and by parental occupation when respondents were 15 years old. Unlike the ACS, the SOEP has only limited information on whether an individual is enrolled in a public or private educational institution. However, the latter play only a small role in Germany.

The traditional three-tiered West German system is comprised of a general secondary school ("Hauptschule," grades 5–9 or 10), an intermediate secondary school ("Realschule," grades 5–10) and an academic high school ("Gymnasium," grades 5–12 or 13). East Germany had a different system with a comprehensive polytechnic secondary school (until grade 10) and an academic track ("Erweiterte Oberschule," grades 11–12) thereafter. The SOEP has detailed information on these different tracks.

We use the SOEP only for the analysis of stylized retrospective education paths of individuals aged 40–45 in 2017. However, as a check on the data, we compare *current* enrollment in the SOEP with the numbers reported by the Federal Statistical Office (*Destatis*). As Table 3 shows, the SOEP matches the official numbers fairly well.

Public Education Expenditure We again use the information on annual public expenditure on education from the OECD, but go back to data from the Federal Statistical Office, which has averages at the state level, as a robustness check. In 2017, the OECD reports \$9.6 K at the primary level, \$13.3 K at the secondary level, and \$10.4 K at the tertiary level (all numbers are in equivalent USD converted using PPPs for GDP). In each case, the expenditure is below the US figures. The difference is most pronounced for tertiary education, where the expenditure in the US is almost three times higher than in Germany.

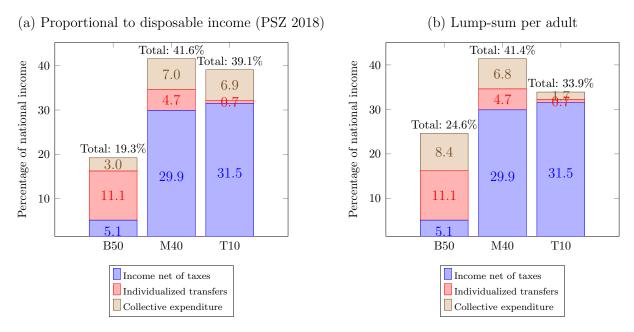
Table 3: Enrollment in Educational Institutions, Germany 2017

	Socio-Economic Panel	Federal Statistical Office
Pre-primary	3.8	3.6
Primary	2.5	2.8
Secondary	4.4	4.5
Tertiary	2.9	2.8
Total	13.6	13.8

Notes: The table compares our SOEP-based numbers for the enrollment in educational institutions in Germany in 2017 with the numbers reported by the Federal Statistical Office. Source: SOEP: Own calculations based on the German Socio-Economic Panel, 2017. Federal Statistical Office: Pre-primary: Statistische Ämter des Bundes und der Länder (2021). Regionaldatenbank Deutschland. Tabelle 22543-01-02-4-B: Betreute Kinder nach Art der Kindertagesbetreuung. https://www.regionalstatistik.de/genesis//online?operation=table&code=22543-01-02-4-B. Primary and secondary: Statistisches Bundesamt (2021). Genesis Online. Tabelle 21111-0003: Schulanfänger: Deutschland, Schuljahr, Geschlecht, Schulart, Einschulungsart. https://www-genesis.destatis.de/genesis//online?operation=table&code=21111-0005: Studierende: Bundesländer, Semester, Nationalität, Geschlecht. https://www-genesis.destatis.de/genesis//online?operation=table&code=21311-0005.

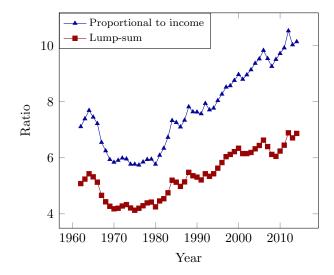
B Additional Tables and Figures

Figure 11: Effects of the Assumptions on the Distribution of Post-tax Income



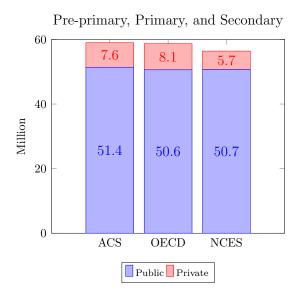
Notes: The figure shows how the assumption regarding the allocation of collective expenditure affects the distribution of post-tax income. When collective expenditure is allocated based on post-tax cash income as in Piketty, Saez, and Zucman (2018), the Bottom 50% receive 19.3% of national post-tax income, while the Middle 40% receive 41.6%, and the Top 10% receive 39.1% (left panel). Under the alternative assumption in which each adult receives the same amount of collective expenditure, the shares are 24.6%, 41.4%, and 33.9% instead (right panel). Source: Own calculations for the United States in 2014 based on Piketty, Saez, and Zucman (2018), Appendix Tables I-SA11, II-C1b, II-C2, II-C3b.

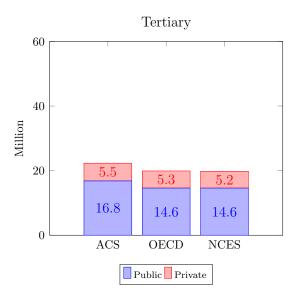
Figure 12: Effect of the Allocation Rules on the Ratio of Average Incomes, 1962–2014



Notes: The figure shows how the assumption regarding the allocation of collective expenditure affects the ratio of average post-tax incomes of the Bottom Top 10% and the Bottom 50% in the United States over the years 1962–2014. The blue graph is for the assumption adopted by Piketty, Saez, and Zucman (2018), i.e. an allocation of collective expenditure that is proportional to post-tax cash income. The red graph shows the ratio that results from assuming a lump-sum allocation. Source: Own calculations based on Piketty, Saez, and Zucman (2018): Appendix Tables I-SA11, II-C1b, II-C2, II-C3b.

Figure 13: Enrollment in Educational Institutions, United States 2017





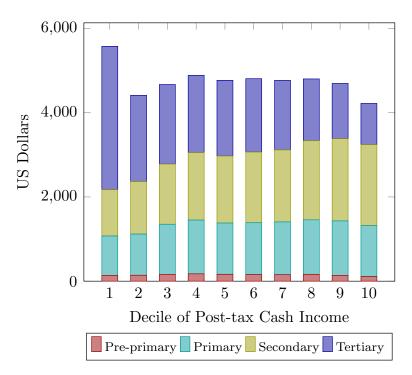
Notes: The figure compares our ACS-based numbers for the enrollment in educational institutions in the United States in 2017 with statistics published by the OECD and the National Center for Education Statistics (NCES). The left panel shows the number of students enrolled in pre-primary, primary, or secondary education, the right panel is for tertiary education. A distinction is made between public and private institutions. Source: Own calculations based on the American Community Survey 2017. OECD: Education at a Glance 2020 (OECD Statistics 2020), Table: Enrollment data adjusted to the financial year. Sum of students in full-time and part-time education. Part-time is only non-zero at the pre-primary and the tertiary levels. Students in post-secondary non-tertiary education not included (110 K are enrolled in public institutions, 273 K in private institutions). NCES: National Center for Education Statistics, Digest of Education Statistics 2019 (De Brey et al. 2021), Table 105.30: Enrollment in elementary, secondary, and degree-granting postsecondary institutions, by level and control of institution: Selected years, 1869-70 through fall 2029.

Table 4: Public Education Spending by Income: Detailed Results

		Income Decile										
	1	2	3	4	5	6	7	8	9	10		
A. Pre-tax Income (Adults)												
Pre-primary	179	165	169	153	146	141	144	152	136	114		
Primary	1,209	1,214	1,230	1,144	1,098	1,118	1,165	1,218	1,245	1,237		
Secondary	1,386	1,447	1,514	1,505	1,504	1,542	1,635	1,758	1,920	1,913		
Tertiary	$3,\!253$	1,927	1,852	1,821	1,817	1,812	1,716	1,534	1,379	992		
Total	6,027	4,752	4,765	4,623	4,564	4,613	4,660	4,661	4,680	$4,\!256$		
B. Post-tax Cash Income (Adults)												
Pre-primary	138	141	162	174	164	157	153	163	138	110		
Primary	935	978	1,189	1,277	1,217	1,231	1,254	1,294	1,293	1,209		
Secondary	1,102	1,248	1,430	1,604	1,591	1,678	1,709	1,877	1,953	1,926		
Tertiary	3,402	2,046	1,889	1,834	1,798	1,742	1,652	1,468	1,312	976		
Total	$5,\!577$	4,413	4,670	4,888	4,769	4,808	4,768	4,802	4,696	4,221		
C. Equivalized Pre-tax Income (Households)												
Pre-primary	379	409	378	333	306	271	241	189	154	124		
Primary	2,701	3,074	2,874	2,627	2,222	2,176	1,923	1,716	1,481	1,290		
Secondary	3,254	3,890	3,890	3,495	3,275	3,152	2,691	2,428	2,049	1,723		
Tertiary	6,771	3,447	3,476	3,524	3,455	3,462	3,147	2,953	2,566	1,686		
Total	13,106	10,820	10,618	9,978	9,258	9,061	8,002	7,286	6,250	4,822		

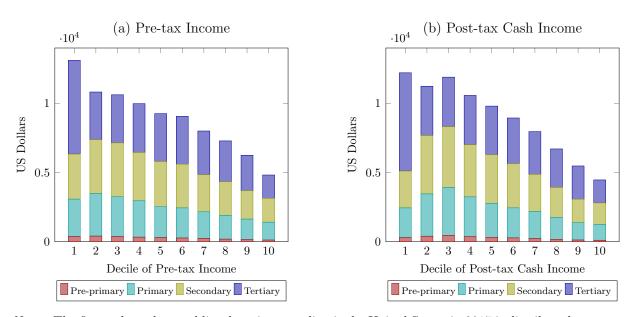
Notes: The table shows how public education spending in the United States in 2017 is distributed among the deciles of the income distribution. All values in 2017 US Dollars. For the sake of presentation and given the large sample size, standard errors are omitted. The deciles are based on pre-tax income (panel A), post-tax cash income (panel B), and equivalized pre-tax household income (panel C). The same information is presented in graphical form in Figure 6 in the main text and Figures 14 and 15 in the Appendix. Source: Enrollment in public educational institutions is taken from the American Community Survey 2017. Each pupil or student is assigned the per-capita value of public education spending taken from the OECD (see Table 2). Public education expenditure is summed up at the household level. In panels A and B, the resulting sum is split equally among adults aged 20 and above in the household, and household income is likewise split equally among all adults. Panel C reports public education spending at the household level instead, and deciles are based on equivalized pre-tax household income, using the modified OECD equivalence scale, which assigns a value of 1 to the first adult in the household, of 0.5 to each additional household member aged 14 and above, and of 0.3 to each child below the age of 14. Pre-tax income is directly taken from the American Community Survey, while post-tax cash income is simulated using TAXSIM v32.

Figure 14: Public Education Spending by Post-tax Cash Income, Allocated Based on Actual Enrollment



Notes: The figure shows how public education spending is the United States in 2017 is distributed among the deciles of the post-tax cash income distribution. For each decile, the bars show the average values of annual public education spending (in 2017 US Dollars) at the pre-primary, primary, secondary, and tertiary levels of education. Source: Enrollment in public educational institutions is taken from the American Community Survey 2017. Each pupil or student is assigned the per-capita value of public education spending taken from the OECD (see Table 2). Public education expenditure is summed up at the household level, and the resulting sum is split equally among adults aged 20 and above in the household. Household post-tax cash income is simulated using TAXSIM v32, and is likewise split equally among all adults.

Figure 15: Public Education Spending by Equivalized Household Income, Allocated Based on Actual Enrollment



Notes: The figure shows how public education spending is the United States in 2017 is distributed among the deciles of the equivalized household income distribution. Left panel: deciles based on pre-tax income. Right panel: deciles based on post-tax cash income simulated using TAXSIM v32. For each decile, the bars show the average values of annual public education spending (in 2017 US Dollars) at the pre-primary, primary, secondary, and tertiary levels of education. Source: Enrollment in public educational institutions is taken from the American Community Survey 2017. Each pupil or student is assigned the per-capita value of public education spending taken from the OECD (see Table 2). Public education expenditure is then summed up at the household level. Household income is equivalized using the modified OECD equivalence scale, which assigns a value of 1 to the first adult in the household, of 0.5 to each additional household member aged 14 and above, and of 0.3 to each child below the age of 14.

Table 5: Robustness Checks

	Decile of Pre-tax Income									
	1	2	3	4	5	6	7	8	9	10
Main specification	6.0	4.8	4.8	4.6	4.6	4.6	4.7	4.7	4.7	4.3
Drop if $> 50\%$ of income imputed	5.9	4.7	4.8	4.6	4.5	4.6	4.7	4.7	4.6	4.2
Drop if income negative	6.0	4.7	4.8	4.6	4.6	4.6	4.7	4.7	4.7	4.3
Drop if income $< 1\%$	6.0	4.7	4.8	4.6	4.6	4.6	4.7	4.7	4.7	4.3
Drop if income $< 2.5\%$	5.4	4.7	4.8	4.6	4.6	4.6	4.7	4.7	4.7	4.3
Drop if income $> 99.5\%$	6.0	4.8	4.8	4.6	4.6	4.6	4.6	4.7	4.7	4.3
Enrollment as in NCES	5.5	4.4	4.4	4.3	4.2	4.3	4.3	4.4	4.4	4.0
Variation across states	5.6	4.4	4.5	4.4	4.3	4.3	4.4	4.5	4.6	4.4
Full-time equivalents 2/4-year college	4.6	3.8	3.8	3.7	3.7	3.7	3.8	3.8	3.9	3.6
	6.8	5.2	5.2	5.1	5.0	5.1	5.1	5.1	5.1	4.6
Current expenditure All expenditure	5.5	4.3	4.3	4.2	4.1	4.2	4.2	4.2	4.2	3.9
	7.2	5.5	5.5	5.4	5.3	5.3	5.4	5.3	5.3	4.8

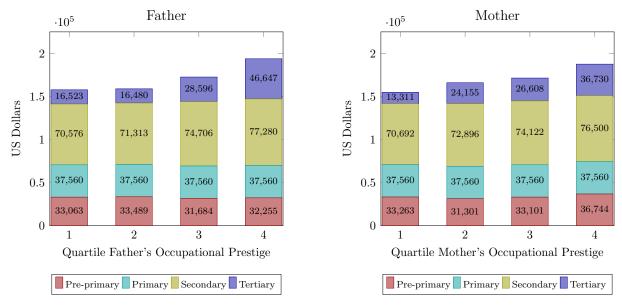
Notes: The table summarizes the results of our robustness checks. For comparison, results for the main specification are shown in the first row as well. The amounts reported in the table are annual public education transfers in thousand US Dollars. Source: In the main specification, enrollment in public educational institutions is taken from the American Community Survey 2017. Each pupil or student is assigned the per-capita value of public education spending taken from the OECD (see Table 2). Public education expenditure is summed up at the household level, and the resulting sum is split equally among adults aged 20 and above in the household. Household income is likewise split equally among all adults. The robustness checks modify the measurement of enrollment, of per-capita expenditure, or of the household income that enters the computation of the deciles. For details, see Section 4.2.

Table 6: Construction of Educational Trajectories

	Years Spent at ISCED Level									
Highest Degree (ACS)	0	1	2	3	4	5	6	7	8	Total
No schooling completed	0	0	0	0	0	0	0	0	0	0
ISCED 0										
Nursery school, preschool	2	0	0	0	0	0	0	0	0	2
ISCED 1										
Kindergarten	2	1	0	0	0	0	0	0	0	3
Grade 1	2	2	0	0	0	0	0	0	0	4
Grade 2	2	3	0	0	0	0	0	0	0	5
Grade 3	2	4	0	0	0	0	0	0	0	6
Grade 4	2	5	0	0	0	0	0	0	0	7
Grade 5	2	6	0	0	0	0	0	0	0	8
ISCED 2										
Grade 6	2	6	1	0	0	0	0	0	0	9
Grade 7	2	6	2	0	0	0	0	0	0	10
Grade 8	2	6	3	0	0	0	0	0	0	11
ISCED 3										
Grade 9	2	6	3	1	0	0	0	0	0	12
Grade 10	2	6	3	2	0	0	0	0	0	13
Grade 11	2	6	3	3	0	0	0	0	0	14
12th grade, no diploma	2	6	3	3	0	0	0	0	0	14
Regular high school diploma	2	6	3	4	0	0	0	0	0	15
GED or alternative credential	2	6	3	4	0	0	0	0	0	15
Some college, but less than 1 year ISCED 4	2	6	3	5	0	0	0	0	0	16
Associate's degree, type not specified ISCED 5	2	6	3	4	2	0	0	0	0	17
1 or more years of college credit, no degree ISCED 6	2	6	3	4	0	2	0	0	0	17
Bachelor's degree	2	6	3	4	0	0	4	0	0	19
ISCED 7	_	Ŭ	J	-		Ŭ	-	Ŭ	Ü	10
Master's degree	2	6	3	4	0	0	4	2	0	21
Professional degree beyond a bachelor's degree	2	6	3	4	0	0	4		0	21
ISCED 8	_	Ü	J	-	Ŭ	V	-	_	V	
Doctoral degree	2	6	3	4	0	0	4	2	4	25

Notes: The table documents how we map the information on the highest degree in the American Community Survey (ACS) 2017 into educational trajectories. The rows correspond to the values of the variable "Highest degree" (educd) in the ACS. The question reads: "What is the highest degree or level of school this person has completed?". As our method is retrospective and we do not have information on grade repetition or, more generally, the individual pathways to a given degree, we assign the same number of years to all individuals with the same degree. For instance, individuals with a regular high school diploma are assumed to have spent two years at ISCED level 0, six years at ISCED level 1, three years at ISCED level 2, and four years at ISCED level 2. Individuals with a bachelor's degree are assigned the same trajectory plus four years at ISCED level 6, and a master's degree would add two years at ISCED level 7. The last column of the table gives the total number of years thus obtained. The number is meant as a summary measure only. When computing the public expenditure for each degree, we multiply the number of years at each ISCED level with the corresponding OECD per-student expenditure from Table 2.

Figure 16: Public Education Spending by Parents' Occupational Prestige. Germany, Individuals Aged 40-45



Notes: The figure depicts average public education spending by parents' occupational prestige for individuals aged 40–45 in Germany. The left panel distinguishes by the occupation of the father, the right panel by the occupation of the mother. Source: Own calculations using the Socio-economic Panel (SOEP) 2017. Public expenditure per student is taken from the "Education at a Glance" database (OECD Statistics 2020). Euro amounts were converted into US Dollars using the GDP purchasing power parities provided by the OECD. Occupational prestige refers to the occupation the parent held when the respondent was 15 years old and is measured using the Standard International Occupational Prestige Score (SIOPS).



Download ZEW Discussion Papers:

https://www.zew.de/en/publications/zew-discussion-papers

or see:

https://www.ssrn.com/link/ZEW-Ctr-Euro-Econ-Research.html https://ideas.repec.org/s/zbw/zewdip.html



ZEW – Leibniz-Zentrum für Europäische Wirtschaftsforschung GmbH Mannheim

ZEW – Leibniz Centre for European Economic Research

L 7,1 · 68161 Mannheim · Germany Phone +49 621 1235-01 info@zew.de · zew.de

Discussion Papers are intended to make results of ZEW research promptly available to other economists in order to encourage discussion and suggestions for revisions. The authors are solely responsible for the contents which do not necessarily represent the opinion of the ZEW.